

**Parental education and occupation in relation to childhood type 1 diabetes:
nationwide cohort study**

Paz Lopez-Doriga Ruiz, MD^{1,2}; German Tapia, PhD¹; Inger J. Bakken, PhD^{3,4}; Siri E. Håberg, MD⁴; Hanne L. Gulseth, MD¹; Torild Skrivarhaug, MD^{5,6}; Geir Joner, MD^{5,6}; Lars C. Stene, PhD¹

¹ Department of Chronic Diseases, Norwegian Institute of Public Health, Oslo, Norway

² Department of Endocrinology, Morbid Obesity and Preventive Medicine, Oslo University Hospital, Oslo, Norway

³ Norwegian Directorate of Health, Trondheim, Norway

⁴ Centre for Fertility and Health, Norwegian Institute of Public Health, Oslo

⁵ Institute of Clinical Medicine, University of Oslo, Oslo, Norway

⁶ Division of Paediatrics and Adolescent Medicine and Oslo Diabetes Research Centre, Oslo University Hospital, Oslo, Norway

Corresponding Author: Lars C. Stene, Department of Chronic Diseases, Norwegian Institute of Public Health, Oslo, Norway, P. O. Box 222 Skøyen, NO-0213 Oslo, Norway. E-mail: lars.christian.stene@fhi.no

Word count: 3068, Abstract: 212

Figures: 3, Tables: 1

References: 36

Supplementary Information: additional text (methods details), 7 Tables, 5 Figures

Background

Socioeconomic status in the risk of developing type 1 diabetes seems inconsistent.

We investigated whether risk of childhood-onset type 1 diabetes differed by parental education or occupation in a nationwide cohort.

Methods

This cohort study included all children born in Norway from 1974 to 2013. In individually linked data from nationwide population registries following children born in Norway up to 15 years of age, we identified 4647 with newly diagnosed type 1 diabetes during 15,381,923 person-years of follow-up.

Results

Children of mothers with a master's degree had lower risk of type 1 diabetes than children of mothers with completed upper secondary education only: adjusted incidence rate ratio, aIRR=0.81 95% confidence interval: 0.69 - 0.95). There was no difference between upper secondary and lower secondary maternal education (aIRR=0.98, 95% confidence interval 0.89-1.08). Paternal education was not significantly associated with type 1 diabetes, lower secondary compared to upper secondary aIRR 0.96 (0.88-1.05) and master compared to upper secondary aIRR 0.93 (0.83-1.05). While maternal elementary occupation was associated with a lower risk of type 1 diabetes, specific maternal- or paternal occupations were not.

Conclusions

Our results suggested inverse U-shaped associations between maternal socioeconomic status and risk of type 1 diabetes. Non-linear associations may be part of the reason why previous literature has been inconsistent.

Background

Type 1 diabetes is among the most common chronic diseases in children and is associated with severe complications and mortality [1]. The increasing incidence among children over the past few decades strongly imply early life environmental factors in the aetiology [2-4]. It is currently not possible to prevent the disease [1], and a better understanding of potentially modifiable risk factors is necessary to identify future targets for primary prevention.

A few environmental risk factors have been weakly but consistently associated with risk of type 1 diabetes, although not established as causal risk factors [4]. Examples include maternal smoking in pregnancy [5, 6], maternal pre-pregnancy obesity [7], childhood infections [8], and short duration of breastfeeding [4]. These potential risk factors tend to be more common in lower socioeconomic groups [9-11].

Another hint that type 1 diabetes may be associated with socioeconomic status is that the incidence of childhood-onset type 1 diabetes is generally higher in wealthier countries [12]. For instance, Norway and other Nordic countries have incidence rates of more than 30 new cases per 100 000 persons per year, while some countries in South America, Africa and Asia have rates below 10 per 100,000 [2]. In fact, the global burden of disease uses economic indices to predict a country incidence of type 1 diabetes [13]. However, this population level trend may in part be due to missed diagnoses in low resource settings and does not imply that an association exists at the individual level.

Our recent literature review identified methodological weaknesses in most previous studies of socioeconomic status in relation to risk of developing type 1 diabetes, including use of ecological- and case-control design and lack of control for confounding by ethnicity or immigration background [14]. Many studies used area-level indicators of socioeconomic status at the time of diagnosis, a design susceptible to-ecological bias. Ecological bias refers to the fact that one cannot generally draw conclusions about individual-level relationships based on group-level data, and the direction of association may even be reversed across levels [15]. For example, even if high income countries tend to have higher population incidence of type 1 diabetes, it is not necessarily children from high socioeconomic status families who tend to develop type 1 diabetes. There is thus a need for high-quality studies of socioeconomic status in relation to risk of type 1 diabetes to assess whether there is an association at the individual level.

Furthermore, some specific occupations involve increased exposures that may be related to the aetiology of type 1 diabetes, independently of socioeconomic status. For example, maternal infections have been linked to type 1 diabetes [16], and frequent social interaction increases the likelihood of many infections [17, 18]. However, few studies have investigated individual level data on maternal or paternal occupation in relation to risk of type 1 diabetes [14]. Examples from other fields of parental occupation and childhood chronic disease risk include childhood cancer risk by parental exposures to chemicals [19]. Farming has been associated with lower risk of asthma and allergies in children. This could be explained by increased exposure to microbes and microbial products leading to a healthier microbiota and lower risk of immune-mediated diseases, according to the hygiene hypothesis [20].

The hygiene hypothesis has also been expanded to include autoimmune diseases such as type 1 diabetes [21]. It has been hypothesized that some infections may increase the risk of type 1 diabetes, while others may reduce the risk, possibly depending on timing and type of infections [4, 16, 21]. Previous studies have reported lower risk of type 1 diabetes in children of parents with occupations involving social contact and thus frequent infections such as teachers and health care workers [17]. Investigating parental occupations may be informative regarding potential environmental exposures in relation to type 1 diabetes.

Nationwide, mandatory registers available in Norway provide a unique opportunity for follow-up for studies of social gradients in incidence of type 1 diabetes in an unselected population. We aimed to examine the potential association of maternal and paternal education and occupation on the incidence of childhood-onset type 1 diabetes in Norway using these linked registers. In addition to assessing socio-economic status, we explored potential association of specific parental occupations.

Methods

Study design and population

This was a cohort study using nationwide registers linked at the individual level by the unique personal identification number assigned to all residents in Norway.

Participants were children born in Norway from 1974 to 2013 followed for type 1 diabetes diagnosed before 15 years of age. We restricted the analyses to children, and parents born in Norway, to ensure completeness and uniformity of data on parental education and occupation (Figure 1). Two sets of linked individual-level data were included, covering two calendar periods with incident cases of type 1 diabetes diagnosed before age 15 years: 1989-2003 and 2005-2013 (Figure 1, details in Supplementary Information). Children born in Norway born during 1974-2013 were followed for type 1 diabetes, and the education and occupation of their parents were obtained from Statistics Norway.

Outcome: Type 1 diabetes

Newly diagnosed, clinical type 1 diabetes was the outcome (Figure 1, details in Supplementary Information).

Parental education

Information on parental education in nine categories was obtained from Statistics Norway, and we used seven categories for the current analyses (see details in Supplementary Information). Because previous high-quality studies had suggested non-linear associations, we decided a priori to use many categories rather than collapsing categories [22, 23]. We used parental education and occupation information collected as close as possible to the time of birth of the child for period 1.

For period 2, maternal and paternal education was only available for 2013, as illustrated in Figure 1a. In 1990, information was only available in a representative sample of the population (of parents). We assessed the effect of analysing information from different time-points in robustness analyses. There has been an increase in the education level in the population over time, and the maternal and paternal education were correlated with a Spearman coefficient of 0.43 (Supplementary Information Figure S1).

Parental occupation

Information on parental occupation was available only in the register linkage for the first study period (Figure 1a). The primary analysis of occupation used 11 categories predefined by Statistics Norway to roughly correspond to socio-economic status. In addition, in an exploratory analysis to assess potential associations of parental occupations with risk of childhood type 1 diabetes, we analysed occupations (two-digit resolution in the Nordic Classification of Occupations) with at least 20 observed cases of type 1 diabetes. Examples of occupations of specific interest were those typically involving social interaction with many children and with infected people. This includes health care workers and teachers. In addition, we were interested in exposure to farm environment.

Covariates

We included the following covariates in adjusted regression models [4]: Maternal and paternal age at delivery, parity, caesarean section, sex, county of residence at birth of the child obtained from the Medical Birth Registry of Norway (age and calendar period was also included as described below). Maternal smoking during pregnancy

was available from the Medical Birth Registry of Norway for births from 1999 (Period 2) and additionally adjusted for in robustness analyses. Maternal type 1 diabetes was available with less than complete coverage for mothers in period 1 and from the Medical Birth Registry for births from 1999 onwards (Period 2). Adjustment for maternal type 1 diabetes was also done in robustness analyses. A conceptual model for how different variables may be linked are shown in Supplementary Information Figure S2.

Statistical analysis

We split individual follow-up times into 1-year categories of attained age and calendar year stratified by exposure and covariates. We estimated adjusted incidence rate ratios (aIRR) with 95% confidence intervals using Poisson regression. We modelled attained age and calendar period flexibly with restricted cubic splines [24, 25] (see supplementary information methods section for more details). For maternal and paternal education, we merged the cross-classified person-time and outcome counts from the two linked datasets. Data were analysed as complete case. The primary statistical test was a test for trend across categories, and we evaluated non-linear associations by categorical analysis and by evaluating a square term for exposure. Analyses were done using STATA version 16 (StataCorp, College Station, TX).

Ethics

The study was approved by Regional Committee for Medical and Health Research Ethics, South-East Norway C (reference 2012/3 data set for study period 1) and

Regional Committee for Medical and Health Research Ethics, South-East Norway B
(reference 2010/2583, data set for study period 2).

Author's Accepted Manuscript

Results

During 15,381,923 person-years of follow-up, 4647 children developed type 1 diabetes (Figure 1). The incidence rate was higher in the most recent period, and slightly higher in boys than in girls (Table 1).

Maternal or paternal education in relation to risk of type 1 diabetes

Maternal education showed an overall significant inverse association with incidence of type 1 diabetes ($p[\text{trend}] = 0.043$), which was non-linear ($p[\text{squared term}] = 0.027$; six degree of freedom likelihood ratio test: $p = 0.011$). There was no difference between level of maternal education among the four lower categories, but a decreasing incidence from the fourth to seventh category (Figure 2). For instance, those with mothers with master's degree had an 18% lower incidence compared to those who only completed upper secondary education (aIRR=0.82 95% confidence interval: 0.70 - 0.95) (Figure 2, further details Supplementary Information Table S1)

Adjustment for covariates gave very similar results to unadjusted analysis (Supplementary Information Figure S3). These results were consistent in period 1 (1989-2004) and 2 (2005-2013) (Supplementary Information Table S2). In robustness analyses, results remained essentially unchanged or even stronger after adjusting for maternal smoking in the sub-cohort with available data (Supplementary Information Figure S4). Characteristics of those with available maternal smoking data were largely similar to that in the whole cohort (Supplementary Information Table S3). Results were also similar after adjusting for maternal type 1 diabetes (Supplementary Information Table S4).

Paternal education was not significantly associated with incidence of type 1 diabetes overall (Figure 2) or in any of the two periods (Supplementary Information Table S2). Robustness analyses showed that associations of maternal or paternal education with type 1 diabetes was similar if we used education at different time points (Supplementary Information Table S5).

Parental occupation and incidence of type 1 diabetes

There was a suggestive non-linear association between maternal occupation categorized according to socioeconomic status, and risk of type 1 diabetes (Figure 3a). Children of mothers with hospitality occupations (for example: hotel, restaurant, household worker, cleaners) had a 23% lower risk of type 1 diabetes compared to children with mothers in professional occupations (aIRR=0.77, 95% CI: 0.63-0.94). Paternal occupation was not significantly associated with type 1 diabetes (Figure 3b). Additional adjustment for parental education gave similar results (Supplementary Information Figure S5). Maternal and paternal occupation was available only for a subset in some of the early censuses, and characteristics were similar for those with and without available data on maternal and paternal occupation (Supplementary Information Table S6).

Finally, we explored the incidence of type 1 diabetes across more detailed occupational categories, including some of specific interest: Farming, which has previously been associated with lower risk of asthma and allergies, and health care workers or teachers who typically exposed to social interaction including people with infectious diseases. However, teachers and farmers ranked average among the common maternal occupations (Supplementary Information Table S7). The

subcategories that were elementary also tended to show lower risk (occupation code 91, 92 and 93 in Supplementary Information Table S7). Hotel, restaurant, and related workers (code 91) ranked lowest in terms of risk for type 1 diabetes among maternal occupations. Subcategories of paternal occupations showed no consistent pattern.

Author's Accepted Manuscript

Discussion

We found that higher maternal education was associated with a slightly lower risk of childhood-onset type 1 diabetes risk, while paternal education was not. Maternal elementary occupations were associated with lower risk of type 1 diabetes, but paternal occupations did not show any consistent association with type 1 diabetes.

Major strengths of our study include the mandatory nationwide register-based design with individual-level information, and large sample size. We accounted for parental country of birth [26], maternal smoking in pregnancy [5], and maternal age [27], potentially important confounders not adjusted for in many previous studies. Notably, since results were remarkably similar before and after adjusting for covariates, these covariates were not likely to mediate or confound much of the observed association with maternal education.

Two other large Scandinavian cohort studies also found inverse U-shaped associations consistent with our result [28, 29]. Two other previous studies, from Italy and USA, respectively, also reported non-linear associations [22, 23]. In the latter studies, the pattern was slightly different compared to our and other Scandinavian findings in that the lowest risk of type 1 diabetes was observed in those with low education. We have no obvious explanation for the different results in the studies from Italy and USA, but it is likely that maternal education in these countries serves as a proxy for a different pattern of unknown exposures here compared to that in Scandinavian countries. The lack of association between paternal education and childhood type 1 diabetes in our study was consistent with that in the only previous cohort study with such data [28].

Our study is to our knowledge the first large scale study to systematically investigate parental occupations in relation to risk of type 1 diabetes. A previous case-control study reported that parental occupations involving exposure to children was associated with lower risk of type 1 diabetes [17]. Consistent with this, we found that the risk of type 1 diabetes tended to be lower in children of mothers in some occupations involving social mixing such as hospitality occupations. This may represent an example of why it is important to study occupational exposures independently of the socioeconomic status associated with an occupation. On the other hand, parental occupational farm exposure, associated with lower risk of asthma and some allergies [30], was not related to lower or higher risk of type 1 diabetes in children in our study. However, our results were consistent with three previous studies investigated farming environment in relation to risk of type 1 diabetes [31-33].

Indicators of socioeconomic status are unlikely to be directly causally related to type 1 diabetes. Interpretation is complex and likely context dependent. However, we think our study contributes important information in a setting where few environmental factors are established as risk factors and several exposures hypothesized to be related to type 1 diabetes have shown to be associated with socioeconomic status. Together with the few previous studies with good methodological quality, our results show that socioeconomic status does not have a very strong relationship with type 1 diabetes. If anything, maternal indicators of socioeconomic status seem more likely relevant than paternal indicators. The underlying suggestive non-linear associations we report may partially explain the inconsistencies among previous studies. This is

because many previous studies had used only two or three categories of socioeconomic status and different (often unreported) cut-offs between categories.

Societal changes during the past four decades include more women with paid jobs, more people with higher education and a higher age at childbirth. We have control for at least some of these factors by adjusting the analyses for year of birth and maternal age. Even though we found consistent associations by period, it is likely that the pattern of unknown exposures that the socioeconomic status indices represent may have changed over the study period. It is certainly possible that unmeasured confounding factors have influenced some of our findings, but we hope our results can inspire others to generate new hypotheses about potential early life exposures to test in future studies.

Our findings also have practical implications for future studies. First, given the relatively weak and inconsistent associations between socioeconomic status and type 1 diabetes, adjustment for socioeconomic status when studying other exposures may not be essential in many settings. Second, we recommend including clearly defined individual-level socioeconomic variables separately for the mother and the father. One of the previous studies with otherwise good methodology used the highest of the father's or mother's education, which complicates interpretation [22]. We speculate that the association with maternal and not paternal education may be driven by intrauterine or early postnatal environmental exposures typically influenced by maternal lifestyle and decisions, such as infant and childhood nutrition and hygiene.

Our study has some limitations. We did not have information on family income. However, our recent review found few other studies and no consistent association between family income and risk of type 1 diabetes [14]. While we focused our analysis on parental education and occupation at the time of the child's birth to avoid potential reverse causality, we did not always have socioeconomic status at the exact year of birth of the child, and data on socioeconomic status and confounders were missing for some periods. On the other hand, our robustness analyses suggested that this did was unlikely to materially influence our results. We cannot generalize our findings to immigrants because information on education and occupation in the registries is incomplete for immigrants. On the other hand, the proportion of immigrants in Norway is too small to have allowed robust analyses in this subgroup even if we had valid information on socioeconomic status. While our large sample size allowed for analysis of parental education and occupation in multiple categories, we decided a priori not to stratify our analyses by sub-groups, because such analyses would have low statistical power and complicate interpretation [34]. We limited our study to children below 15 years, and our results cannot automatically be extrapolated to older ages. Finally, it may be argued that Nordic countries have a relatively egalitarian society with universal and free access to health care and education, and that limited social inequality in health is to be expected. However, clear social gradients in adult mortality and self-rated health have been found in the Nordic countries [35]. Furthermore, higher maternal education has been robustly associated with lower childhood adiposity [11] and lower childhood mortality in the Nordic countries [36]. We therefore believe that differences in socioeconomic status within Norway are not negligible and that we should be able to detect associations with our large sample size, if there truly is a relationship.

Conclusions

In conclusion, our results suggested associations between maternal socioeconomic status, with a slightly lower risk of type 1 diabetes in children of mothers with higher education. Paternal socioeconomic status was not associated with type 1 diabetes in children.

Author's Accepted Manuscript

Availability of data. Norwegian data protection legislation and Act on medical and health research do not allow individual level patient data to be shared by the authors. However, all data are accessible to authorised researchers after ethical approval and application to the registries via <https://helsedata.no/> administered by the Directorate of eHealth, and the Norwegian Childhood Diabetes Registry.

Duality of Interest. PLDR reports participation in a research project funded by LEO Pharma, all regulator mandated phase IV studies, all with funds paid to her institution (no personal fees) and with no relation to the work reported in this paper. All other authors declare no competing interests.

Funding. This research was funded in part by a grant from the South-Eastern Norway Regional Health Authority and by the Norwegian Institute of Public Health, and in part by the Research Council of Norway through its Centers of Excellence funding scheme, project number 262700.

Author Contributions. LCS and GJ conceptualized the study and research question. GJ, LCS and SEH were responsible for data acquisition, planning, and funding. LCS and PLDR conceptualized the detailed study design. PLDR, GT and IJB carried out the data preparation. PLDR and LCS did the statistical analyses with input and quality control by GT. All authors contributed to the interpretation of results. PLDR wrote the first draft of the manuscript and subsequent revisions. All authors critically revised the paper for important intellectual content and approved the final version. PLDR and LCS had full access to all the data in the study and take responsibility for the integrity of the data and the accuracy of the data analysis.

“What is already known on this topic”

- Incidence of type 1 diabetes is higher in wealthier countries, with highest incidence in Nordic countries and lower incidence in South American and Asian countries.
- Individual-level epidemiological studies of socioeconomic status in relation to risk of type 1 diabetes in children seems sparse and heterogeneous in methodology and results.

“What this study adds”

- Maternal education showed a threshold association with lower risk among children of mothers with higher university education.
- Paternal education was not associated with type 1 diabetes.
- Maternal elementary occupations were associated with a lower risk of type 1 diabetes.

“How this study might affect research, practice or policy”

- Non-linear associations may help explained why previous literature has been inconsistent.

References

1. DiMeglio, L.A., C. Evans-Molina, and R.A. Oram, *Type 1 diabetes*. Lancet, 2018. **391**(10138): p. 2449-2462.
2. Tuomilehto, J., et al., *Update on worldwide trends in occurrence of childhood type 1 diabetes in 2020*. *Pediatr Endocrinol Rev*, 2020. **17**(Suppl 1): p. 198-209.
3. Rewers, M. and J. Ludvigsson, *Environmental risk factors for type 1 diabetes*. Lancet, 2016. **387**(10035): p. 2340-2348.
4. Norris, J.M., R.K. Johnson, and L.C. Stene, *Type 1 diabetes-early life origins and changing epidemiology*. *Lancet Diabetes Endocrinol*, 2020. **8**(3): p. 226-238.
5. Magnus, M.C., et al., *Parental smoking and risk of childhood-onset type 1 diabetes*. *Epidemiology*, 2018. **29**(6): p. 848-856.
6. Edstorp, J., A.M. Lampousi, and S. Carlsson, *Parental smoking, type 1 diabetes, and islet autoantibody positivity in the offspring: A systematic review and meta-analysis*. *Diabet Med*, 2022: p. e14830.
7. Magnus, M.C., et al., *Paternal and maternal obesity but not gestational weight gain is associated with type 1 diabetes*. *Int J Epidemiol*, 2018. **47**(2): p. 417-426.
8. Isaacs, S.R., et al., *Enteroviruses and risk of islet autoimmunity or type 1 diabetes: systematic review and meta-analysis of controlled observational studies detecting viral nucleic acids and proteins*. *Lancet Diabetes Endocrinol*, 2023.
9. Dowd, J.B., A.E. Aiello, and D.E. Alley, *Socioeconomic disparities in the seroprevalence of cytomegalovirus infection in the US population: NHANES III*. *Epidemiol Infect*, 2009. **137**(1): p. 58-65.
10. Lund-Blix, N.A., et al., *Infant feeding and risk of type 1 diabetes in two large Scandinavian birth cohorts*. *Diabetes Care*, 2017. **40**(7): p. 920-927.
11. Biehl, A., et al., *Adiposity among children in Norway by urbanity and maternal education: a nationally representative study*. *BMC Public Health*, 2013. **13**: p. 842.
12. Gomez-Lopera, N., N. Pineda-Trujillo, and P.A. Diaz-Valencia, *Correlating the global increase in type 1 diabetes incidence across age groups with national economic prosperity: A systematic review*. *World J Diabetes*, 2019. **10**(12): p. 560-580.
13. GBD 2021 Diabetes Collaborators, *Global, regional, and national burden of diabetes from 1990 to 2021, with projections of prevalence to 2050: a systematic analysis for the Global Burden of Disease Study 2021*. Lancet, 2023. **402**(10397): p. 203-234.
14. Lopez-Doriga Ruiz, P. and L.C. Stene, *Is socio-economic status associated with risk of childhood type 1 diabetes? Literature review*. *Diabet Med*, 2023: p. e15182.
15. Morgenstern, H. and J. Wakefield, *Ecological studies and analysis*, in *Modern epidemiology*, T.L. Lash, et al., Editors. 2021, Wolters Kluwer: Philadelphia.
16. Allen, D.W., et al., *Maternal virus infections in pregnancy and type 1 diabetes in their offspring: Systematic review and meta-analysis of observational studies*. *Rev Med Virol*, 2018. **28**(3): p. e1974.
17. Ponsonby, A.L., et al., *Higher parental occupational social contact is associated with a reduced risk of incident pediatric type 1 diabetes: Mediation through molecular enteroviral indices*. *PLoS One*, 2018. **13**(4): p. e0193992.
18. Magnusson, K., et al., *Occupational risk of COVID-19 in the first versus second epidemic wave in Norway, 2020*. *Euro Surveill*, 2021. **26**(40).
19. Savitz, D.A. and J.H. Chen, *Parental occupation and childhood cancer: review of epidemiologic studies*. *Environ Health Perspect*, 1990. **88**: p. 325-37.
20. von Mutius, E. and D. Vercelli, *Farm living: effects on childhood asthma and allergy*. *Nat Rev Immunol*, 2010. **10**(12): p. 861-8.
21. Bach, J.F., *The hygiene hypothesis in autoimmunity: the role of pathogens and commensals*. *Nat Rev Immunol*, 2018. **18**(2): p. 105-120.

22. Bruno, G., et al., *Early life socioeconomic indicators and risk of type 1 diabetes in children and young adults*. J Pediatr, 2013. **162**(3): p. 600-605 e1.
23. D'Angeli, M.A., et al., *Environmental factors associated with childhood-onset type 1 diabetes mellitus: an exploration of the hygiene and overload hypotheses*. Arch Pediatr Adolesc Med, 2010. **164**(8): p. 732-8.
24. Carstensen, B., *Age-period-cohort models for the Lexis diagram*. Stat Med, 2007. **26**(15): p. 3018-45.
25. Royston, P. and P.C. Lambert, *Flexible parametric survival analysis using stata: beyond the Cox model*. 2011, College Station, Texas: Stata Press.
26. Dzidzonu, D.K., et al., *Ethnic differences in the incidence of type 1 diabetes in Norway: a register-based study using data from the period 2002-2009*. Pediatr Diabetes, 2016. **17**(5): p. 337-41.
27. Cardwell, C.R., et al., *Maternal age at birth and childhood type 1 diabetes: a pooled analysis of 30 observational studies*. Diabetes, 2010. **59**(2): p. 486-494.
28. Clausen, T.D., et al., *Broad-spectrum antibiotic treatment and subsequent childhood type 1 diabetes: a nationwide Danish cohort study*. PLoS One, 2016. **11**(8): p. e0161654.
29. Khashan, A.S., et al., *Gestational age and birth weight and the risk of childhood type 1 diabetes: a population-based cohort and sibling design study*. Diabetes Care, 2015. **38**(12): p. 2308-15.
30. von Mutius, E., *The "hygiene hypothesis" and the lessons learnt from farm studies*. Front Immunol, 2021. **12**: p. 635522.
31. Radon, K., et al., *Exposure to farming environments in early life and type 1 diabetes: a case-control study*. Diabetes, 2005. **54**(11): p. 3212-3216.
32. Heikkinen, S.M., et al., *Does farm environment protect against type 1 diabetes mellitus?* Diab Vasc Dis Res, 2013. **10**(4): p. 375-7.
33. Virtanen, S.M., et al., *Microbial exposure in infancy and subsequent appearance of type 1 diabetes mellitus-associated autoantibodies: a cohort study*. JAMA Pediatr, 2014. **168**(8): p. 755-763.
34. Oxman, A.D. and G.H. Guyatt, *A consumer's guide to subgroup analyses*. Ann Intern Med, 1992. **116**(1): p. 78-84.
35. Mackenbach, J.P., et al., *Socioeconomic inequalities in health in 22 European countries*. N Engl J Med, 2008. **358**(23): p. 2468-81.
36. Gissler, M., et al., *Sex differences in child and adolescent mortality by parental education in the Nordic countries*. J Epidemiol Community Health, 2012. **66**(1): p. 57-63.

Table 1. Characteristics of cohort giving rise to incident cases of type 1 diabetes during two periods. Children born in Norway to Norwegian-born parents.

		Period 1989-2003			Period 2005-2013		
		Person-years †	T1D cases	Incidence rate*	Person-years†	T1D cases	Incidence rate*
	Total	9192165	2325	25.3 (24.3 - 26.3)	6292314	2338	37.2 (35.7 - 38.7)
Child's sex	Boys	4723688	1242	26.3 (24.9 - 27.8)	3228802	1260	39.0 (36.9 - 41.2)
	Girls	4468300	1083	24.2 (22.8 - 25.7)	3063513	1078	35.2 (33.1 - 37.4)
Maternal age at delivery	<20	269577	79	29.3 (23.5 - 36.5)	163294	61	37.4 (29.1 - 48.0)
	20-24	2261783	544	24.1 (22.1 - 26.2)	1013357	388	38.3 (34.7 - 42.3)
	25-29	3661699	915	25.0 (23.4 - 26.7)	2187115	835	38.2 (35.7 - 40.9)
	30-34	2272315	589	25.9 (23.9 - 28.1)	1997983	723	36.2 (33.6 - 38.9)
	≥ 35	726791	198	27.2 (23.7 - 31.3)	930538	331	35.6 (31.9 - 39.6)
Parity (birth order)	1st	3889582	990	25.5 (23.9 - 27.1)	2568947	909	35.4 (33.2 - 37.8)
	2nd	3419412	870	25.4 (23.8 - 27.2)	2282775	861	37.7 (35.3 - 40.3)
	3rd	1456259	348	23.9 (21.5 - 26.5)	1063028	421	39.6 (36.0 - 43.6)
	4th	329224	93	28.2 (23.1 - 34.6)	276206	115	41.6 (34.7 - 50.0)
	≥5th	97688	24	24.6 (16.5 - 36.7)	101358	32	31.6 (22.3 - 44.6)
Caesarean section	No	8146319	2041	25.1 (24.0 - 26.2)	5379958	1974	36.7 (35.1 - 38.3)
	Yes	1045846	284	27.2 (24.2 - 30.5)	912357	364	39.9 (36.0 - 44.2)
Paternal age at delivery	<25	1192733	315	26.4 (23.6 - 29.5)	568355	214	37.7 (32.9 - 43.1)
	25-29	3288251	810	24.6 (23.0 - 26.4)	1690550	661	39.1 (36.2 - 42.2)
	30-34	3046489	780	25.6 (23.9 - 27.5)	2170022	780	35.9 (33.5 - 38.6)
	35-39	1313297	348	26.5 (23.9 - 29.4)	1253151	481	38.4 (35.1 - 42.0)
	≥40	351394	72	20.5 (16.3 - 25.8)	573854	193	33.6 (29.2 - 38.7)

* Rate per 100,000 person-years (95% confidence interval). T1D: Type 1 diabetes. T1D cases: Number of incident cases during follow-up.

† Total cohort size was 1,139,859 children in period 1989-2003 and 1,125,142 children in period 1990-2013. Many children contributed with person-years in both periods (See also Figure 1a).

Figure 1. Illustration of study design and formation of analysis sample

a) Birth cohorts, age and timing of parental education and education, and type 1 diabetes (T1D) in children. Incident T1D before age 15 years was ascertained by the Norwegian Childhood Diabetes Registry during 1989-2003 (red box) and by incident use of insulin at least twice according to the Norwegian Prescription Database during 2005-2014 (green box). b) Formation of the analysis sample. * Two sets of linked data were split in one-year intervals of attained age and calendar year. Person-time and incident cases in each age- and calendar year combination were merged to obtain independent contributions from each of the two linked data sets in the merged data for analysis. Detailed explanation in Supplemental Material, p.2. † Missing parental education (both parents).

Figure 2. Parental education and incidence of type 1 diabetes (T1D)

a) Maternal education and b) Paternal education. Nationwide cohort of children in Norway of whom 4647 developed T1D before 15 years of age during 1989-2013. Incidence rate ratios (aIRR) were adjusted for parental age at delivery, parity, caesarean section, child's sex, county of residence, age and calendar period. Vertical lines represent 95% confidence intervals for the aIRRs. Statistical test for trend: $p=0.043$ for maternal education and $p=0.83$ for paternal education. There was significant deviation from linearity for maternal education (square term $p=0.027$; six degree of freedom likelihood ratio test for maternal education: $p=0.011$). Details regarding numbers in each category, absolute incidence rates and characteristics for each category of maternal education are shown in Supplemental Table S1.

Figure 3. Parental occupation and incidence of type 1 diabetes (T1D)

a) Maternal occupation, and b) paternal occupation. Nationwide cohort of children in Norway of whom 1,497 developed T1D before 15 years of age during 1989-2003. Incidence rate ratios (aIRR) were adjusted for parental age at delivery, parity, caesarean section, child's sex, county of residence, age and calendar period. Vertical lines represent 95% confidence intervals (CI) for the aIRRs. The aIRR for mothers who worked in mining and quarry industry was not shown in the figure due to low number of cases, but it was 5.65, 95% CI: 1.41-22.7. Likelihood ratio test for overall association (10 degree of freedom): $p=0.017$ for maternal occupation and $p=0.64$ for paternal occupation. Additional adjustment for maternal education gave similar results (Figure S4).