

The development of a nonstationary standardised streamflow index using climate and reservoir indices as covariates

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1 Abstract

2 Under current global change, the driving force of evolution of drought has
3 gradually transitioned from a single natural factor to a combination of natural and
4 anthropogenic factors. Therefore, widely used standardised drought indices based on
5 assumption of stationarity are challenged and may not accurately assess characteristics
6 of drought processes. In this study, a nonstationary standardised streamflow index
7 (NSSI) that incorporates climate and reservoir indices as external covariates was
8 developed to access nonstationary hydrological drought. The first step of the proposed
9 approach is to apply methods of trend and change point analysis to assess the
10 nonstationarity of streamflow series to determine type of streamflow regime, that is, the

11 natural and altered regime. Then, different nonstationary models were constructed to
12 calculate the NSSI by selecting climate indices as covariates for streamflow series with
13 natural regime, and climate and reservoir indices as covariate for streamflow series with
14 altered regime. Four stations in the upper reaches of the Huaihe River basin, China,
15 were selected to examine the performance of the proposed NSSI. The results indicated
16 that Dapoling (DPL), Changtaiguan (CTG), and Xixian (XX) stations had natural
17 streamflow regimes, while the Nanwan (NW) station had an altered regime. The global
18 deviances of the optimal nonstationary models were 17 (2.2%), 18 (2.9%), 26 (4.0%),
19 and 22 (3.5%) less than those of stationary models for DPL, CTG, NW, and XX stations,
20 respectively. Especially, for the NW station influenced by reservoir regulations, the
21 frequency of slight drought and moderate drought of NSSI was 12.8% lower than and
22 13.1% greater than those of SSI, respectively. Overall, the NSSI that incorporates the
23 influence of climate variability and reservoir regulations provided more reliable
24 assessment of hydrological drought than the traditional SSI.

25

26 **Keywords:** Hydrological drought; Nonstationary standardised streamflow index;
27 Climate variability; Reservoir regulation

28

29 **1. Introduction**

30 Droughts are regarded as complex and multidimensional phenomena that pose a
31 threat to water security worldwide ([Araghinejad 2011](#); [Ali et al. 2021](#); [Wang et al. 2021](#)).
32 Traditionally, droughts are classified into four categories, including meteorological,
33 agricultural, hydrological, and socioeconomic drought ([David and Davidov 2017](#);
34 [AghaKouchak et al. 2021](#)). As an important form of drought, hydrological drought can
35 cover extensive areas and can last for months to years ([Van Loon 2015](#)). For example,

36 in 2014, the state of California in the USA faced one of the most severe multiyear
37 droughts on record, resulting in extremely low reservoir and groundwater levels
38 (Aghakouchak et al. 2014). Therefore, accurate understanding and quantification of
39 hydrological droughts are essential for water management and policy making.

40 Several hydrological drought indices have been proposed over the past few
41 decades and they are generally divided into two groups—standardised indices and
42 threshold-based indices. The most widely used standardised indices are the
43 standardised runoff index (SRI; Shukla and Wood 2008) and the standardised
44 streamflow index (SSI; Vicente-Serrano et al. 2012). Both indices have a calculation
45 procedure similar to standardised precipitation index (SPI), which involves fitting a
46 probability distribution to a hydrological variable and transforming the distribution to
47 a normal distribution. Another method for deriving hydrological drought events is the
48 threshold level method (Van loon and Van Lanen 2013). This approach sets a pre-
49 defined threshold level, and when the target variable (e.g., streamflow) is below this
50 level, the site is considered to be in a state of drought.

51 Traditionally, drought indices are calculated based on the fundamental assumption
52 of statistical stationarity and assumed that the probabilistic characteristics of the hydro-
53 meteorological process do not change over time. However, Milly et al (2008) argued
54 that the effects of climate change and anthropogenic activities challenge the
55 fundamental assumption of stationarity. Thus, in a changing environment, alternative
56 approaches should be developed to consider nonstationarity and allow probabilistic
57 parameters to change over time (Salas and Obeysekeru 2014). In this way, statistical
58 distribution parameters are expressed as functions of covariates to model
59 nonstationarity generated by climatic and anthropogenic impacts.

60 Several studies have attempted to consider the impact of nonstationarity on the

61 evaluation of hydrological droughts. [Zou et al \(2018\)](#) used the generalised additive
62 model for location, scale, and shape (GAMLSS) to construct a time-dependent SRI
63 (SRI_{var}), with its parameters linked with time to fit a long-term streamflow time series,
64 to assess the impacts of climate change and human activities on hydrological drought.
65 However, many studies have indicated that nonstationary models in which the
66 parameters vary only with time cannot adequately capture the variability of hydro-
67 meteorological variables ([Wang et al. 2021c](#)). In recent years, many researchers have
68 emphasised that choosing large-scale climate indices as covariates of streamflow in
69 nonstationary models would better characterise the nonstationarity caused by climate
70 variability ([Forootan et al. 2019](#)). In addition to climate variability, human activity is
71 another driver of nonstationarity ([Jiang et al. 2019](#)). Among diverse anthropogenic
72 activities (e.g., irrigation, dams and reservoirs, and inter-basin water transfer), reservoir
73 construction and regulation usually cause the greatest shifts in flow regimes ([Slater et](#)
74 [al. 2021](#)). Some recent studies have attempted to parameterise the influence of reservoir
75 construction and regulation to investigate the resulting nonstationary impacts on
76 streamflow processes. For example, [López and Francés \(2013\)](#) carried out
77 nonstationary flood frequency analysis for continental Spanish rivers using climate and
78 reservoir indices (RI) as external covariates. [Jiang et al. \(2015\)](#) selected time and RI as
79 covariates to explore how reservoirs altered low-flows. However, research on
80 nonstationary hydrological drought assessment with the incorporation of climate
81 variability and anthropogenic reservoir regulations remains limited.

82 Here, we develop a nonstationary standardised streamflow index (NSSI) using the
83 GAMLSS, with its parameters vary nonlinearly with climate indices and RI. The
84 specific objectives of the study were to (1) use trend and change point analysis methods
85 to reveal the nonstationarity of streamflow series under the impacts of climate

86 variability and reservoir regulations; (2) select the most significant climate indices and
87 calculate the RI for use as covariates to construct the NSSI based on the GAMLSS
88 model; and (3) compare the NSSI with the traditional SSI in the evaluation of
89 hydrological drought to assess the efficiency of the NSSI. The Xixian basin, located in
90 the upper reaches of the Huaihe River Basin, China, was chosen as a case study to
91 evaluate the developed NSSI. The NSSI is expected to provide a useful tool for the
92 evaluation of nonstationary hydrological drought in a changing environment.

93

94 **2. Study area and data**

95 2.1 Study area

96 The Xixian basin, located in the upper reaches of the Huaihe River Basin in eastern
97 China (31.5°N–32.75°N, 113.25°E–115°E), covers an area of 10,191 km² (Fig. 1). The
98 basin is situated in the transition zone of mid-latitude humid and semi-arid climates
99 (Jiang et al. 2020), and both flood and hydrological drought events occur frequently.
100 We selected Dapoling, (DPL), Changtaiguan (CTG), Nanwan (NW), and Xixian (XX)
101 stations as the case study areas. Several studies have indicated that three large reservoirs
102 (i.e., Nanwan, Feishahe, and Huashan reservoirs in Fig. 1) have impacts on hydrological
103 extremes (e.g., hydrological drought) in the study area (Shi et al. 2011).

104

105 2.2 Data

106 The observed monthly precipitation of the 35 rain-gauge stations and streamflow
107 of the four hydrological stations during 1955–2016, and information of three large
108 reservoirs in the upper reaches of NW station were obtained from the Hydrological
109 Bureau of the Ministry of Water Resources of China. Capacity and completion date of
110 the Nanwan, Feishahe, and Huashan reservoirs were 1630, 82, 173 million m³ and 1955,

111 1961, 1967, respectively. According to previous studies (Ouyang et al. 2014), six large-
112 scale climate indices with proven influence on the streamflow of Xixian basin, i.e., AO,
113 AMO, NAO, PDO, NPO, and ENSO were selected as candidate covariates. They were
114 obtained from the National Oceanic and Atmospheric Administration (NOAA)
115 Physical Sciences Laboratory website (<https://psl.noaa.gov/data/climateindices/>).

116

117 3. Methods

118 The framework of the proposed NSSI consisted of three steps as shown in Fig. 2.
119 (1) Nonstationary analysis of hydro-meteorological variables; (2) Construction of the
120 NSSI; and (3) Validation of the NSSI. The methods used in this study are described in
121 detail in following sections.

122

123 3.1 Nonstationary analysis methods

124 The robustness of the traditional Mann-Kendall (MK) trend test is usually
125 influenced by the persistence of the hydro-meteorological series. To improve the MK
126 test method, Hamed and Rao (1998) used *lag-i* autocorrelation to remove persistence
127 and make the test more robust, known as the modified Mann-Kendall (MMK) trend test.
128 In this study, we used the MMK method to identify trends of the hydro-meteorological
129 series.

130 Conventional statistical test methods for identification of change points such as
131 Mann-Kendell test and sliding T test are usually based on the assumption that the time
132 series should be linear and stationary. However, it is difficult for them to accurately
133 capture the change points in the nonlinear or nonstationary time series. To make up this
134 shortcoming, this study applied the heuristic segmentation method, proposed by
135 Bernaola-Galvan et al. (2001), to identify change points in nonlinear and nonstationary

136 series. In addition, the precipitation-streamflow double cumulative curve (DCC)
137 method was also applied to identify change points. Generally, a change in the gradient
138 of the curve implies that the characteristics of precipitation or streamflow may have
139 changed, and the inflection point of the curve is generally regarded as a change point
140 (Wang et al. 2020b).

141

142 3.2 Calculation of the NSSI

143 A 3-month SSI (SSI-3) will not cause the SSI value to change too fast or too slowly.
144 Therefore, SSI-3 (Vicente-Serrano et al. 2012) was considered more suitable for the
145 identification of continuous drought events in the Xixian basin. In this study, the
146 drought grades can be divided into five classes based on the SRI values: non-drought
147 when $SRI > -0.5$, slight drought (D1) when $-1.0 < SRI \leq -0.5$, moderate drought (D2)
148 when $-1.5 < SRI \leq -1.0$, severe drought (D3) when $-2.0 < SRI \leq -1.5$, and extreme
149 drought (D4) when $SRI \leq -2.0$. Besides, a stationary model (M1) and a nonstationary
150 model (M2) were developed to calculate SSI and NSSI, respectively. In M1, a
151 streamflow series was fitted to a stationary model with constant distribution parameters
152 to construct a stationary SSI series. In M2, a streamflow series was fitted to a
153 nonstationary model based on the GAMLSS model with distribution parameters
154 varying with climate indices and/or the RI to construct a NSSI series. The main steps
155 to construct the NSSI includes: (1) selection of significant climate indices and
156 calculation of RI; and (2) non-stationary probability fitting of streamflow based on
157 GAMLSS model and calculation of the NSSI.

158

159 3.2.1 Selection of significant climate indices

160 To reduce the monthly random variability (Wang et al. 2020c) and maintain

161 consistency with NSSI-3 (or SSI-3), each climate index was processed using a 3-month
 162 moving average. With regard to the streamflow series in a specified month, the
 163 reconstructed time series considering lead times of 1-, 2-, ..., 12-months were prepared
 164 for the correlation analysis. Kendall's rank correlation test was performed at the 95%
 165 confidence level to identify significant climate indices that can be considered covariates
 166 to fit the streamflow records in the nonstationary model.

167

168 3.2.2 Calculation of the RI

169 The dimensionless RI proposed by López and Francés (2013) is an effective
 170 indicator to reflect the impact of reservoir construction on flow regimes. It is calculated
 171 as follows:

$$172 \quad RI = \sum_{i=1}^N \left(\frac{A_i}{A_T} \right) \cdot \left(\frac{C_i}{C_T} \right) \quad (1)$$

173 where N is the number of reservoirs upstream of the gauge station, A_i is the catchment
 174 area of each reservoir, A_T is the catchment area of the gauge station, C_i is the total
 175 capacity of each reservoir, and C_T is the mean annual streamflow at the gauge station.
 176 The RI threshold value between low and high alteration was found to be 0.25, and the
 177 higher the RI value, the greater the reservoir influence.

178

179 3.2.3 GAMLSS

180 The GAMLSS proposed by Rigby and Stasinopoulos (2005) assumed that
 181 independent observations x_i for $i = 1, 2, 3, \dots, n$ have the distribution function $f(x_i | \theta^i)$
 182 where $\theta^i = (\theta_1^i, \theta_2^i, \dots, \theta_p^i)$ is a vector of p distribution parameters accounting for the
 183 location, scale, and shape random variable characteristics. The distribution parameters
 184 are related to covariates by the monotonic link functions $g_k(\cdot)$ for $k = 1, 2, \dots, p$, where

185 the parameters are modelled through proper link functions:

$$186 \quad g_k(\theta_k) = \eta_k = X_k \beta_k + \sum_{j=1}^{J_k} Z_{jk} x_{jk} \quad (2)$$

187 where η_k and θ_k are vectors of length n , for example $\theta_k^T = (\theta_k^1, \theta_k^2, \dots, \theta_k^n)$; X_k is a matrix
 188 of explanatory variables (i.e., covariates) of order $n \times J_k$; $\beta_k^T = (\beta_{1k}, \beta_{2k}, \dots, \beta_{J_k k})$ is a
 189 parameter vector of length J_k ; Z_{jk} is a fixed known $n \times q_{jk}$ design matrix; and x_{jk} is a q_{jk} -
 190 dimensional random variable.

191 In this study, we selected six widely used two-parameter continuous distributions
 192 to fit the streamflow data, including the Weibull, Gumbel, Gamma, Logistic, Normal,
 193 and Lognormal distributions (Vicente-Serrano et al. 2012; Wang et al. 2020a). The
 194 model that achieved the minimum Global deviance (GD), Akaike information criterion
 195 (AIC), and Schwarz Bayesian information criterion (BIC) values was selected as the
 196 optimal nonstationary model. A visual inspection of the diagnostic plots of the residuals
 197 (i.e., worm plot) was also used to ensure that the selected model adequately described
 198 the systematic part of each time series. All calculations were performed using the R
 199 platform and the freely available ‘*gamlss*’ package. Mathematically, the calculation of
 200 the NSSI is similar to that of the SSI but is based on the selected optimal nonstationary
 201 models. The classification of drought grades for the NSSI is the same as that of the SSI.

202

203 3.3 Copula and joint return period

204 In this study, we selected the “and (\cap)” joint return period (Salvadori and De
 205 Michele 2004) to calculate the joint probability (P_{DS}) and joint return period (T_{DS}) of
 206 drought duration and severity:

$$207 \quad P_{DS} = P(D \geq d, S \geq s) = 1 - F_D(d) - F_S(s) + C(F_D(d), F_S(s)) \quad (3)$$

$$208 \quad T_{DS} = \frac{E(L)}{P(D \geq d \cap S \geq s)} = \frac{E(L)}{1 - F_D(d) - F_S(s) + C(F_D(d), F_S(s))} \quad (4)$$

209 where $C(F_D(d), F_S(s))$ is the joint distribution of drought duration and severity, and $E(L)$
 210 represents the average time interval of drought events.

211 The six continuous distributions in [section 3.2.3](#) were selected to determine the
 212 optimal marginal distribution of the drought severity series. Two widely used discrete
 213 distributions namely the Negative Binomial and Poisson distributions were applied to
 214 fit the drought duration series. The distribution that showed the smallest d value, as
 215 calculated by the Kolmogorov-Smirnov (K-S) test, was selected as the optimal marginal
 216 distribution. Then, the optimal copula function was selected from five commonly used
 217 theoretical copula functions, namely the Gaussian-copula, T-copula, Clayton-copula,
 218 Frank-copula, and Gumbel-copula functions ([Wang et al. 2020a](#)).

219

220 4. Results

221 4.1 Nonstationary analysis of hydro-meteorological series

222 [Figs. 3\(a\)-\(b\)](#) show that the seasonal distributions of precipitation of the four
 223 stations are consistent, with most of the precipitation falling between May and
 224 September. Streamflow of the NW station peaks in and concentrated around June, while
 225 that of the other stations peak in and concentrated around July. DCCs of the [Figs. 3\(c\)-](#)
 226 [\(f\)](#) illustrate that precipitation and streamflow of the DPL, CTG, and XX stations
 227 remained consistent, whereas this relationship of the NW station had been disturbed
 228 since 1956. [Table 1](#) shows that streamflow of the NW station showed a significant
 229 downward trend ($\alpha < 0.05$). Heuristic segmentation test identified two significant change
 230 points ($\alpha < 0.05$) of the NW station in 1956 and 1967, respectively. According to these
 231 results, the DPL, CTG, and XX stations can be considered as stations with a ‘natural’
 232 streamflow regime and the NW station can be selected as a station with a ‘altered’

233 streamflow regime.

234

235 4.2 Construction of the NSSI

236 [Table 2](#) shows the identification results of significant climate indices in CTG
237 station. The results indicated that ENSO was selected as the significant index in January
238 to February and August to December; NPO was regarded as a crucial impact factor for
239 streamflow in March, April, June, and October to December; PDO dominantly affected
240 streamflow in August to November; AMO was selected as the significant climate index
241 in all the months. Besides, the maximum RI values (i.e., nowadays) were 0.0014,
242 0.0012, 0.3203, and 0.0049 for DPL, CTG, NW, and XX stations, respectively. The RI
243 value of the NW station was larger than the threshold (i.e., 0.25), while that of the other
244 stations was far less than 0.25. [Figs. 3](#) (g), (h), and (j) show that for stations with a low
245 RI, reservoir construction had less impact on the streamflow regime. However, for the
246 NW station with a high RI value ([Fig. 3](#) (i)), the streamflow regime has been strongly
247 altered since 1967. Therefore, for stations with natural regimes, significant climatic
248 indices should be considered when building nonstationary models. For the NW station,
249 alongside climatic indices, the RI will also be considered as an external covariate.

250 [Table 3](#) shows that the all the GD, AIC, and BIC values of M2 were smaller than
251 those of M1 in all the 12 months at NW station. The GD of M2 was reduced by 9 (1.2%
252 in July) to 28 (4.4% in April) compared with M1. [Table 4](#) shows that the Lognormal
253 distribution (LOGNO) gave the best fits for the three natural stations, and the Logistic
254 distribution (LO) gave the best fit for the altered NW station. The average GDs of M2
255 model were 17 (2.2%), 18 (2.9%), 26 (4.0%), and 22 (3.5%) less than those of M1
256 model for DPL, CTG, NW, and XX stations, respectively.

257 [Fig. 4](#) shows that M2 performed better than M1 when describing the variability of

258 the streamflow series in normal (March), wet (June), and dry (December) months. For
259 example, in normal month (Fig. 4(a)), the area spanned by the 5th and 95th quantiles
260 of M1 covered 94%, 90%, 89% and 94% of the observed points at the four stations.
261 Whereas that of M2 covered 98%, 98%, 94% and 97% of the observed points,
262 respectively. Fig. 5 shows that all residual points of M1 and M2 for the three natural
263 stations fell within the 95% confidence intervals. However, the residual values of the
264 NW station for M1 (Fig. 5 (c)) shows deviations from normality, while those of M2
265 (Fig. 5 (g)) do not show a departure from normality. Therefore, the nonstationary M2
266 that incorporated climate indices and RI as covariates can evidently provide a better fit
267 for streamflow series. Then, the optimal M2 models were used to fit the streamflow
268 data and to calculate the NSSI series.

269

270 4.3 Validation of the NSSI

271 According to the historical records (Chinese Office of State Flood Control and
272 Drought Relief Headquarters, OSFCDRH 1997), two historical drought events that
273 occurred in 1978–1979 and 1993 at the NW station were chosen to examine the
274 performance of the NSSI and SSI. The extreme droughts 1978–1979 (the first gray
275 shadow bars in Fig. 6(a)) caused a serious reduction in crop production in the Xixian
276 basin. The NSSI captured an extreme hydrological drought event from August 1978 to
277 April 1980, which reached the most (i.e., -2.65) in September 1979. Whereas, the SSI
278 values during this period were all below the level of severe drought (i.e., $SSI > -2$). In
279 1993, a hydrological drought event occurred from spring to autumn in NW station (the
280 second gray shadow bars in Fig. 6(a)), among which drought in autumn had the greatest
281 impact (Zhu et al. 2016). NSSI index captured the gradual aggravation trend of this
282 hydrological drought event from spring to autumn, while the drought process captured

283 by SSI had been maintained at the level of moderate drought (i.e., $SSI > -1.5$).

284 [Fig. 6\(d\)](#) shows that the frequency of slight (D1) and moderate (D2) drought of
285 NSSI of the NW station were 12.8% lower and 13.1% greater than that of SSI,
286 indicating that reservoir regulations in the upper reach of the NW station caused more
287 slight droughts developed into moderate droughts. Besides, Negative Binomial (NB)
288 and LOGNO distributions were the optimal marginal distribution functions for drought
289 duration and severity of the four stations, respectively. Frank-copula function was the
290 optimal copula functions for the joint distribution of drought duration and severity of
291 these stations. [Figs. 6\(f\)-\(m\)](#) show that joint contour plots of the NSSI (M2) were
292 moving forward in comparison with those calculated from the SSI (M1), indicating that
293 the drought risk identified by the NSSI series is increased compared with that of SSI
294 series.

295

296 **5. Discussion**

297 The traditional SSI is widely used for hydrological drought assessment and it is
298 generated by fitting streamflow series to a distribution with constant parameters based
299 on stationary assumption ([Wang et al., 2020c](#)). However, effects of climate variability
300 and anthropogenic activities challenge the fundamental assumption of stationarity, and
301 statistical distribution parameters may change with these forcing factors ([Jiang et al.](#)
302 [2015](#)). In this study, the proposed NSSI were generated from probability models that
303 distribution parameters varying with significant climate indices and RI, thus providing
304 an effective tool for nonstationary hydrological drought assessment. On the one hand,
305 comparisons of criteria values, i.e., the GD, AIC, and BIC between M1 and M2 ([Tables](#)
306 [3 and 4](#)) showed that nonstationary models had smaller fitting residuals. Centile curves
307 and worm plots ([Figs.4 and 5](#)) showed that model M2 provided a better description of

308 streamflow processes than M1. On the other hand, the verification results (Fig. 6)
309 showed that the NSSI that incorporate the climate variability and reservoir regulations
310 will give a more reliable description of drought condition and process.

311 Climate variability is widely recognised as one of the most important multi-years
312 to multi-decadal drivers of changes in hydro-climatic extremes (Slater et al. 2021). As
313 shown in Fig. 4, the nonstationary models considering different lead times when
314 selecting the climate indices were better at describing streamflow and drought
315 conditions. Such attempts will contribute to a deeper understanding of historical
316 hydrological drought events at catchment and regional scales (Van Loon 2015). More
317 importantly, the advantage of employing climate variability is that climate model
318 predictions can be employed as covariates to estimate future changes in drought
319 processes (Slater et al. 2021). In addition, with the increase in global population and
320 the growing demand for water, anthropogenic impacts on surface hydrological
321 processes have greatly increased (AghaKouchak et al. 2014, 2021). Among diverse
322 anthropogenic water use activities, reservoir regulations usually have a more direct
323 influence on flow regimes (Slater et al. 2021). The Figs. 4-6 revealed that for the NW
324 station, the nonstationary models considering the RI as a covariate provided a more
325 accurate fit of the streamflow data, especially changes occurring before and after
326 streamflow change points related to reservoir construction. Such attempts provide an
327 idea for parameterizing the human influence and then applying it as covariates to
328 construct a nonstationary standard streamflow index.

329 Finally, some limitations should be emphasized is that the proposed NSSI in this
330 study only considered the nonstationarity caused by climate variability and reservoir
331 regulation. When the NSSI is applied to other regions, other nonstationary forcing
332 factors should be considered, such as land use and cover change, irrigation, and

333 abstraction (AghaKouchak et al. 2021; Slater et al. 2021). The future work could focus
334 on developing some parameterisation schemes that can account for human activities
335 and then applying them to the NSSI construction framework.

336

337 **6. Conclusions**

338 In this study, we developed a NSSI that incorporates large-scale climate indices
339 and the RI as external covariates based on the GAMLSS model. The large-scale climate
340 indices were selected based on correlation analysis to represent impacts of climate
341 variability, and the RI was calculated to represent influence of anthropogenic reservoir
342 regulation. As such, our comprehensive consideration of the nonstationary impacts of
343 these factors on hydrological drought assessment is the innovative part of this research.

344 We demonstrated the application of the developed NSSI to the Xixian basin
345 located in the upper reaches of the Huaihe River Basin, China. For this basin, the
346 average GDs of M2 model were 17 (2.2%), 18 (2.9%), 26 (4.0%), and 22 (3.5%) less
347 than those of M1 model for DPL, CTG, NW, and XX stations, respectively, indicating
348 that the nonstationary models performed better than stationary model in describing the
349 nonstationary streamflow processes. Especially, for the NW station, the frequency of
350 slight drought and moderate drought of NSSI was 12.8% lower than and 13.1% greater
351 than those of SSI, respectively. These revealed that the NSSI can correctly capture the
352 variations of hydrological drought due to influence of reservoir regulations.

353 Overall, the developed NSSI in this study provides an appropriate method for
354 assessing nonstationary hydrological droughts in a changing environment. Such
355 attempts are expected to provide a valuable reference for nonstationary drought
356 research in other regions as well as the development of new drought assessment
357 methods.

Author Contribution

Conceptualization: Menghao Wang; Methodology: Linyong Wei, Hao Cui, Fei Yuan, Yi Liu, and Xiaoli Yang; Funding acquisition: Shanhu Jiang, Liliang Ren, and Chong-Yu Xu.

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Data Availability

The data that support the findings of this study are openly available at <https://www.usgs.gov/> and <https://psl.noaa.gov/data/climateindices/>.

Declarations**Ethical Approval**

Not applicable.

Consent to Participate

Not applicable.

Consent to Publication

Not applicable.

Competing Interests

The authors declare that they have no competing interests.

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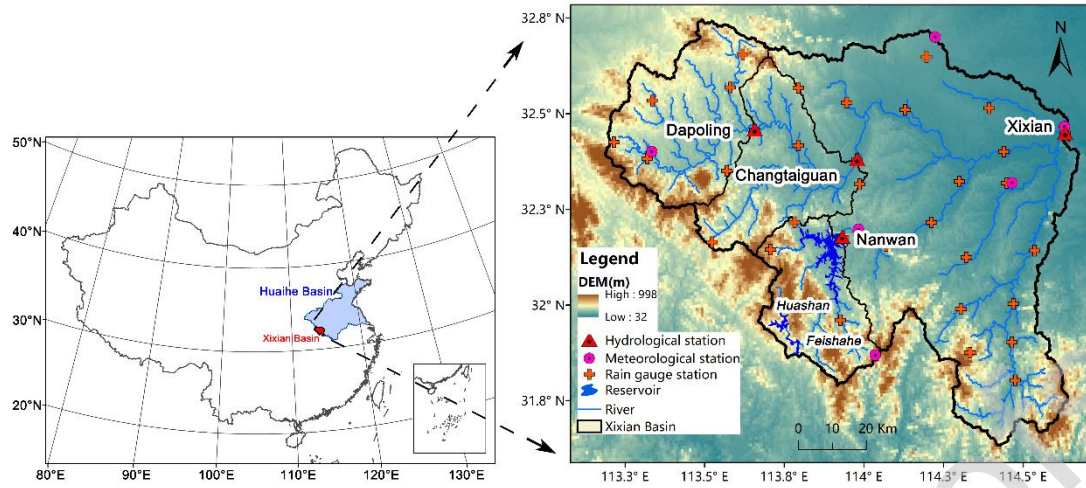


Fig. 1. Location of the study area and the distribution of hydrological and rain gauge stations and reservoirs.

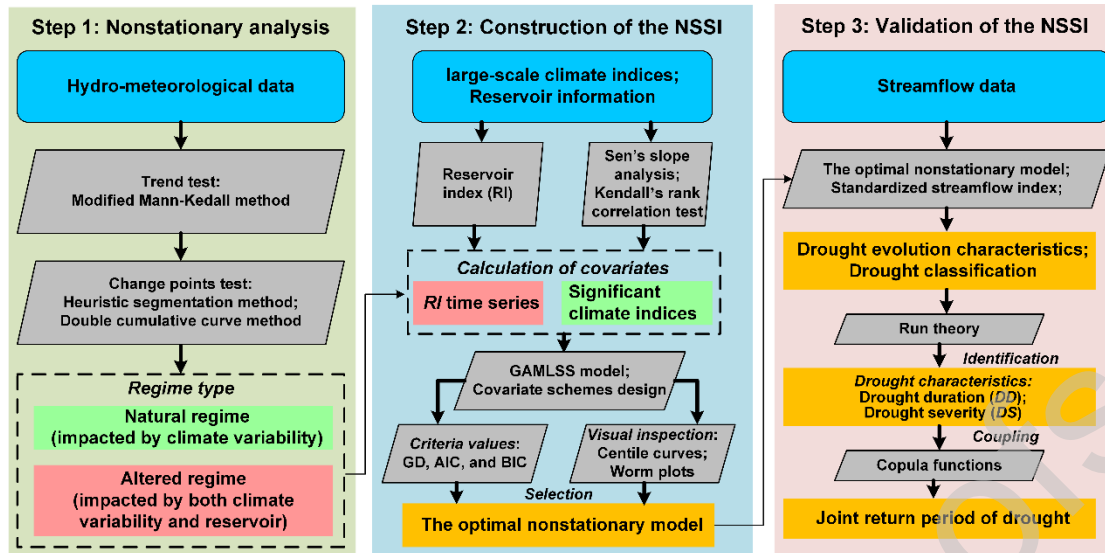


Fig. 2. Framework for the construction and validation of the nonstationary standardised streamflow index (NSSI) proposed in this study.

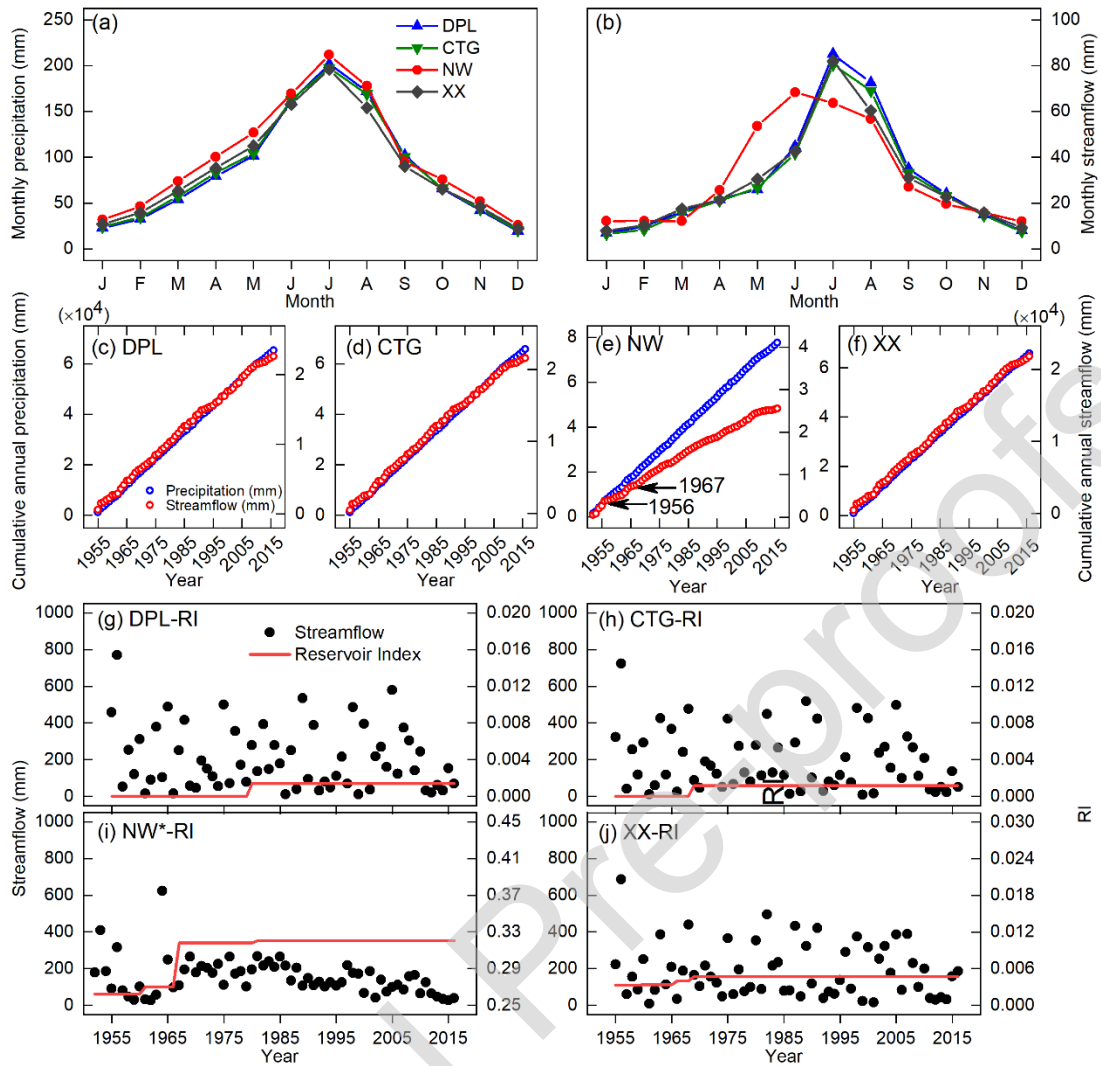


Fig. 3. Monthly (a) precipitation and (b) streamflow series; (c)-(f) Double cumulative curves of annual precipitation and streamflow for the four stations; (g)-(j) Temporal evolution of the reservoir index (RI) and 3-month cumulative streamflow series in June for the four stations (* indicates a highly altered streamflow regime).

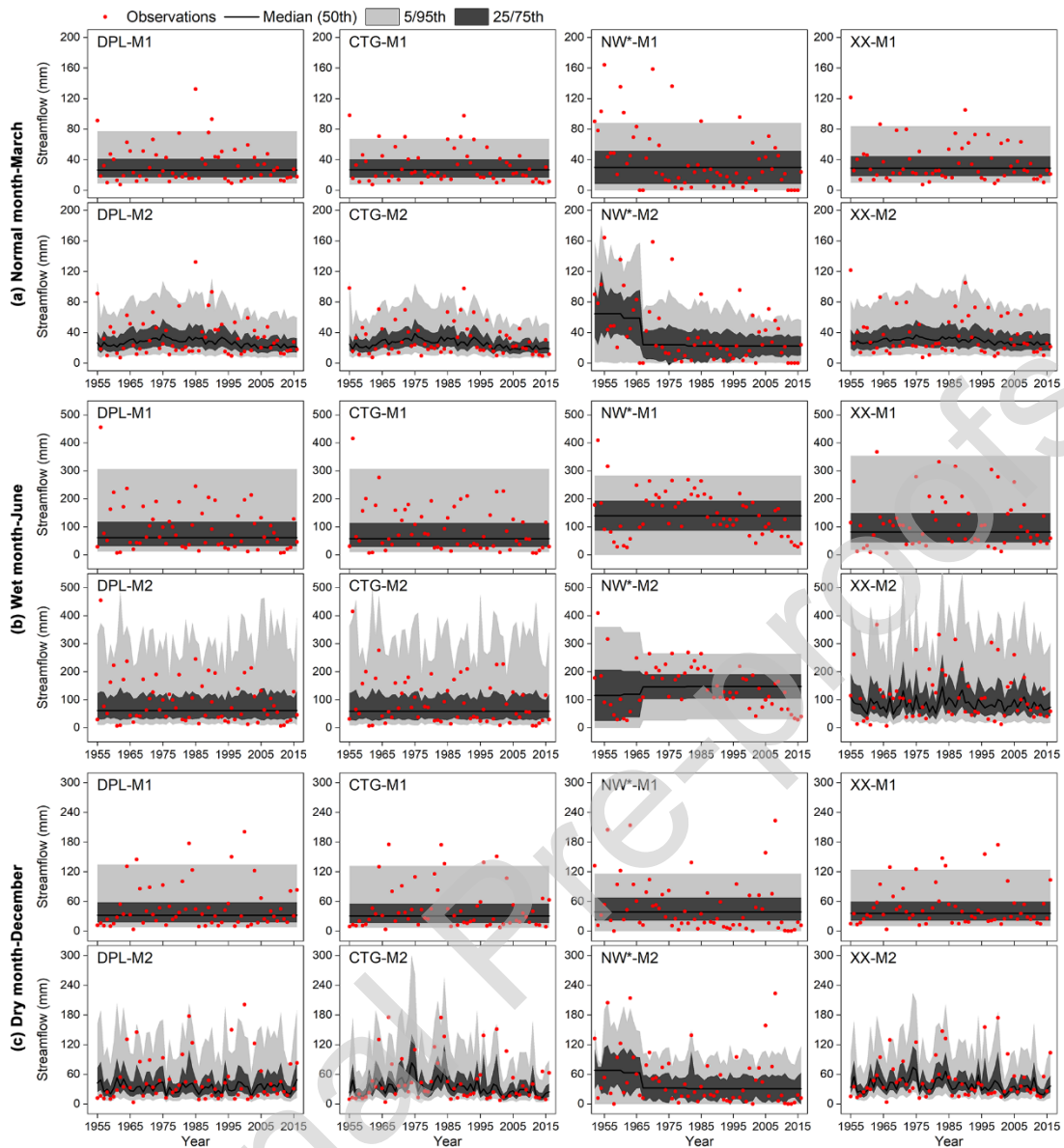


Fig. 4. Centile curves plots for modelling 3-month cumulative streamflow at the four stations with a stationary model (M1) and nonstationary model (M2) during (a) normal month (March), (b) wet month (June), and (c) dry month (December). ‘*’ indicates a highly altered flow regime. The red points are the observed streamflow series; the blank line is the median (50% centile curve); the dark-grey region represents the area between the 25% and 75% centile curves; and the light-grey region represents the area between the 5% and 95% centile curves.

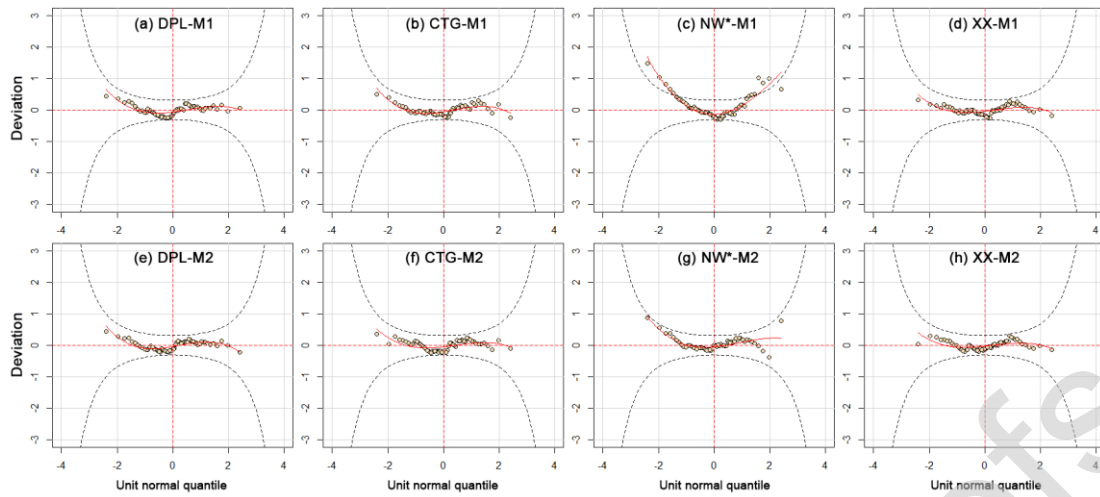


Fig. 5. Worm plots of residuals of the optimal M1 and M2 models for observed streamflow (3-month cumulative streamflow series in March) at the four stations. ‘*’ indicates a highly altered flow regime. The two black dotted lines correspond to the 95% confidence limits. For a good fit, the data points should be aligned preferably along the red solid line but within the two dashed black lines.

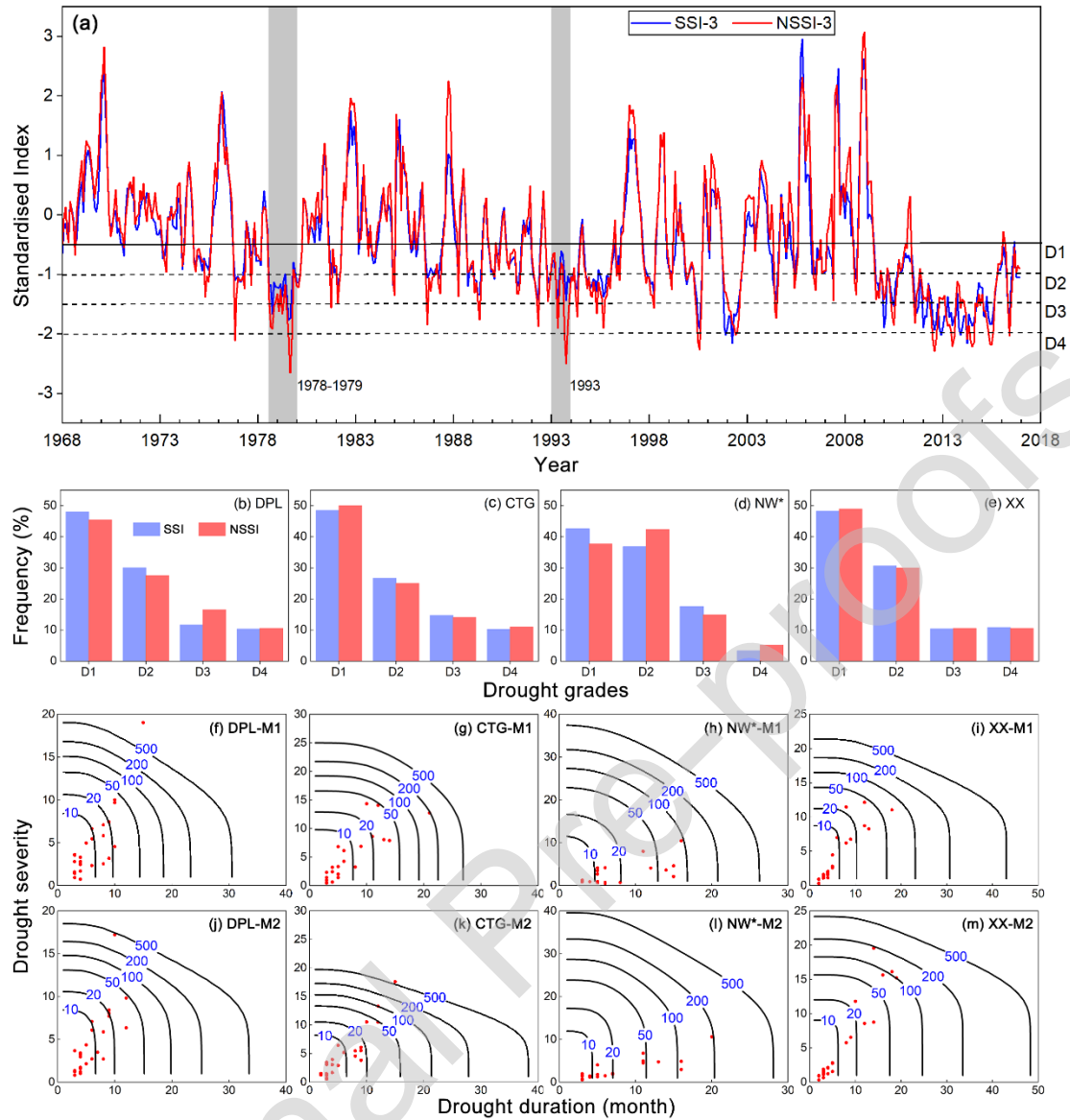


Fig. 6. (a) Temporal series of the standardised streamflow index (SSI) and the nonstationary standardised streamflow index (NSSI) over a 3-month scale at the NW station (the gray shadow bars indicate the hydrological drought events of 1978–1979 and 1993, respectively); (b)-(e) Frequencies of different drought grades (D1–D4) obtained from SSI-3 and NSSI-3 of the four stations; (f)-(m) Joint return periods of hydrological drought duration and severity obtained from the SSI-3 (i.e., M1) and the NSSI-3 (i.e., M2) for the four stations. ‘*’ indicates a highly altered flow regime and red dots indicate the observed hydrological drought events.

Table 1. Results of the trend and change point test of annual precipitation and streamflow for the four stations in the Xixian Basin.

Station	MMK trend test (Z values)		Heuristic segmentation test for change point	
	Precipitation	Streamflow	Precipitation	Streamflow
DPL	-0.92↓	-1.58↓	—	—
CTG	-0.69↓	-1.63↓	—	—
NW	-0.26↓	-2.16* ↓	—	1956*, 1967*
XX	-0.54↓	-1.48↓	—	—

Notes: ‘*’ with bold values denotes significance at the 95% confidence level. ‘↑’ and ‘↓’ indicate upward and downward trends, respectively.

Table 2. Summary of the significant climate indices for each month of a year at the CTG station.

Month	Climate index						Month	Climate index					
Jan	<i>ENSO₁₁</i>	NAO ₆	AO ₆	NPO ₁	PDO ₁₁	<i>AMO₁</i>	Jul	ENSO ₆	<i>NAO₁₁</i>	<i>AO₅</i>	NPO ₅	PDO ₅	<i>AMO₇</i>
	0.18	-0.12	-0.14	-0.09	0.13	-0.28		0.13	0.18	-0.21	0.16	0.12	-0.23
Feb	<i>ENSO₁₂</i>	NAO ₇	AO ₁₂	NPO ₅	PDO ₁₂	<i>AMO₂</i>	Aug	<i>ENSO₇</i>	NAO ₈	AO ₆	NPO ₆	<i>PDO₆</i>	<i>AMO₈</i>
	0.18	-0.12	-0.12	-0.10	0.13	-0.29		0.19	0.14	-0.14	0.14	0.18	-0.24
Mar	ENSO ₃	NAO ₈	AO ₁	<i>NPO₇</i>	PDO ₁₁	<i>AMO₄</i>	Sep	<i>ENSO₈</i>	NAO ₆	AO ₇	<i>NPO₇</i>	<i>PDO₇</i>	<i>AMO₉</i>
	-0.14	-0.10	0.10	-0.18	-0.11	-0.29		0.20	0.16	-0.11	0.11	0.23	-0.25
Apr	ENSO ₄	NAO ₈	AO ₉	<i>NPO₇</i>	PDO ₀	<i>AMO₅</i>	Oct	<i>ENSO₈</i>	NAO ₁₀	<i>AO₀</i>	<i>NPO₄</i>	<i>PDO₈</i>	<i>AMO₁₀</i>
	-0.11	-0.14	-0.09	-0.22	-0.14	-0.28		0.20	0.10	0.20	0.21	0.24	-0.27
May	ENSO ₅	NAO ₂	AO ₁	NPO ₅	PDO ₁	<i>AMO₆</i>	Nov	<i>ENSO₉</i>	NAO ₁₂	AO ₄	<i>NPO₉</i>	<i>PDO₉</i>	<i>AMO₁₁</i>
	-0.10	0.13	0.09	-0.14	-0.16	-0.23		0.18	-0.10	-0.10	0.17	0.19	-0.24
Jun	ENSO ₁₂	NAO ₉	<i>AO₄</i>	<i>NPO₆</i>	PDO ₀	<i>AMO₆</i>	Dec	<i>ENSO₁₀</i>	NAO ₃	AO ₅	<i>NPO₀</i>	PDO ₁₀	<i>AMO₀</i>
	-0.11	0.13	-0.18	-0.18	-0.14	-0.21		0.19	-0.12	-0.09	-0.18	0.13	-0.24

Notes: Subscript numbers represent the lead times between each climate index and the corresponding month's streamflow. The numbers listed below each climate index are the Kendall correlation coefficients. The climate index with correlation coefficient greater than 95% confidence level is expressed in bold and italics.

Table 3. Summary of the assessment criteria values (GD, AIC, and BIC) for the optimal M1 and M2 models at the NW station and its selected covariate for μ and σ in the optimal M2 model. RI refers to reservoir index and ct refers to a parameter that is independent of the covariates (i.e., constant).

Month	GD		AIC		BIC		The optimal M2 model	
	M1	M2	M1	M2	M1	M2	μ	σ
Jan	619	609	623	617	627	606	RI	RI
Feb	596	585	600	595	604	486	RI	RI, PDO ₁₁
Mar	622	604	626	614	630	604	RI	RI, AMO ₄
Apr	649	622	653	636	658	631	RI, AMO ₅	RI, AMO ₅ , NPO ₈
May	703	678	707	688	712	688	AMO ₆	RI, AMO ₆
Jun	729	720	733	728	738	717	RI	RI
Jul	750	741	754	749	758	727	RI	RI
Aug	742	732	746	740	750	729	RI	RI
Sep	739	716	743	724	747	722	ct	RI, AMO ₁₀
Oct	719	698	723	708	727	698	ct	RI, AO ₀ , AMO ₁₁
Nov	699	680	703	690	707	681	ct	RI, NPO ₀ , AMO ₁₂
Dec	658	646	662	656	666	636	RI	RI, NPO ₀

Table 4. Summary of the optimal M1 and M2 models for the four hydrological stations.

Station	Distribution		GD		AIC		BIC	
	M1	M2	M1	M2	M1	M2	M1	M2
DPL	LOGNO	LOGNO	641	627	645	623	650	629
CTG	LOGNO	LOGNO	637	619	641	626	645	623
NW	LO	LO	685	659	689	669	694	669
XX	LOGNO	LOGNO	636	614	640	619	644	624

Notes: LOGNO and LO represent the Lognormal and Logistic distributions, respectively.